

## 18.600 Midterm 2, Fall 2019 Solutions

1. (20 points)

- (a) Melissa is applying to 20 different out-of-state medical schools. Because of her excellent GPA/MCAT/essays, her chance of being accepted to each school is  $1/20$ , and the decisions at the 20 schools are independent of each other. Using a Poisson approximation, estimate the probability that Melissa will be accepted to at least two of these schools. **ANSWER:** Number  $X$  of acceptances is roughly Poisson with parameter  $\lambda = 20 \cdot \frac{1}{20} = 1$ . Thus  $P(X \geq 2) = 1 - P(X = 1) - P(X = 0) \approx 1 - e^{-\lambda}\lambda^1/1! - e^{-\lambda}\lambda^0/0! = 1 - 2/e \approx .26424$ . **Remark:** If we compute the exact value using a binomial distribution, we get  $P(X \geq 2) \approx .26416$ , so the approximation is quite good.

- (b) Jill is applying to 25 different out-of-state medical schools and has a  $1/5$  chance (independently) of being invited for an interview at each school. Let  $X$  be the number of medical schools at which she is invited to interview. Compute  $E[X]$  and  $\text{Var}[X]$ . **ANSWER:** The number of interviews is binomial with parameter  $n = 25$  and  $p = 1/5$ . So  $E[X] = np = 5$  and  $\text{Var}[X] = np(1 - p) = 4$ .

- (c) Using a normal approximation, roughly approximate the probability that Jill is invited to interview at fewer than 2.5 schools. You may use the function

$$\Phi(a) = \int_{-\infty}^a \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$

in your answer. **ANSWER:** Since the standard deviation of  $X$  is 2, the value 2.5 is  $5/4$  standard deviations below the mean. Hence the probability is approximately  $\Phi(-5/4) \approx .10565$ . **Remark:** The true probability is .098 which is pretty close.

2. (20 points) A room has four lightbulbs, each of which will burn out at a random time. Let  $X_1, X_2, X_3, X_4$  be the burnout times, and assume they are independent exponential random variables with parameter  $\lambda = 1$ . Write

1.  $X = X_1 + X_2 + X_3 + X_4$ .
2.  $Y = \min\{X_1, X_2, X_3, X_4\}$ , i.e.,  $Y$  is time when first bulb burns out.
3.  $Z = \max\{X_1, X_2, X_3, X_4\}$ , i.e.,  $Z$  is time when last bulb burns out.

Compute the following:

- (a) The probability density function  $f_X$ . **ANSWER:** This is a Gamma distribution with parameters  $\lambda = 1$  and  $n = 4$ . So  $f_X(x) = x^3 e^{-x}/3!$  for  $x \in [0, \infty)$ .
- (b) The probability density function  $f_Y$ . **ANSWER:** The minimum of four exponentials of parameter 1 is exponential with parameter 4. Hence  $f_Y(x) = 4e^{-4x}$  for  $x \in [0, \infty)$ .

(c) The expectation  $E[Z]$ . **ANSWER:** This is basically the radioactive decay problem from lecture. Answer is  $1/4 + 1/3 + 1/2 + 1$ .

(d) The covariance  $\text{Cov}(Y, Z)$ . (Hint: use memoryless property.) **ANSWER:** The memoryless property implies that  $Y$  and  $Z - Y$  are independent and hence  $\text{Cov}(Y, Z) = \text{Cov}(Y, Y + (Z - Y)) = \text{Cov}(Y, Y) = \text{Var}(Y)$ . Since  $Y$  is exponential with parameter  $\lambda = 4$  its variance is  $1/\lambda^2 = 1/16$ .

3. (20 points) Five applicants are applying for a job, and an interviewer gives each applicant a score between 0 and 1. Call these scores  $X_1, X_2, \dots, X_5$  and assume that they are i.i.d. uniform random variables on  $[0, 1]$ . The top applicant has score  $Y = \max\{X_1, X_2, \dots, X_5\}$ , and the second to the top has score  $Z$ , which we define to be the *second* largest of the  $X_i$ . Compute the following:

(a) The cumulative distribution function  $F_Y(r)$  for  $r \in [0, 1]$ . **ANSWER:**  
 $P(Y \leq r) = P(\max\{X_1, X_2, \dots, X_5\} \leq r) = P(X_1 \leq r, X_2 \leq r, \dots) = P(X_1 \leq r)^5 = r^5$ .

(b) The density function  $f_Y$ . **ANSWER:**  $f_Y(r) = F'_Y(r) = 5r^4$  for  $r \in [0, 1]$  (and zero if  $r \notin [0, 1]$ ).

(c) The density function  $f_Z$  and the value  $E[Z]$ . **NOTE:** If you remember what this means, you may use the fact that a Beta  $(a, b)$  random variable has expectation  $a/(a + b)$  and density  $x^{a-1}(1 - x)^{b-1}/B(a, b)$ , where  $B(a, b) = (a - 1)!(b - 1)!/(a + b - 1)!$ . **ANSWER:** The ordering of candidates is independent of the set of scores obtained by the candidates. This means that the density of  $Z$  is the same that of a uniform random variable conditioned on three people being smaller, one being larger. This is a Beta  $(a, b)$  random variable with  $a - 1 = 3$  and  $b - 1 = 1$ . So it comes to  $x^3(1 - x)/B(4, 2) = 20x^3(1 - x)$  and  $E[X] = 4/(4 + 2) = 2/3$ .

(d) The probability  $P(X_2 > 2X_1)$  (i.e., probability second candidate's score is more than than double first candidate's score). **ANSWER:** Note that joint density  $f_{X_1, X_2}(x, y)$  is 1 on the unit square  $[0, 1]^2$  and zero elsewhere. Therefore the probability is the area of the subset of  $[0, 1]^2$  where  $y > 2x$ , which comes to  $1/4$ . So the answer is  $1/4$ .

4. (15 points) Let  $X$  and  $Y$  be independent random variables with density function given by  $\frac{1}{\pi(1+x^2)}$ .

(a) Compute  $P(X < 1)$ . **ANSWER:**  $X$  is a Cauchy random variable, so the answer is  $3/4$  by our spinning flashlight story. Recall that in that story, we draw a line from  $(0, 1)$  with a uniformly chosen angle and its intersection with  $\mathbb{R}$  is a Cauchy random variable. The angle range corresponding to  $(-\infty, 1)$  is  $3/4$  of the total range, so the answer is  $3/4$ .

(b) Compute the probability density function for the random variable  $Z = (X - Y)/2$ . **ANSWER:** If  $Y$  is Cauchy then  $-Y$  is also Cauchy. The average of two independent Cauchy random variables is itself Cauchy, so the answer is  $\frac{1}{\pi(1+x^2)}$ .

(c) Compute  $E[e^{-X^2-Y^2}]$ . You can leave your answer as a double integral—no need to evaluate it explicitly. **ANSWER:**  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{1}{\pi(1+x^2)} \frac{1}{\pi(1+y^2)} e^{-x^2-y^2} dx dy$

5. (10 points) Let  $X_1, X_2, X_3, \dots, X_{10}$  be the outcomes of independent standard die rolls—so each takes one of the values in  $\{1, 2, 3, 4, 5, 6\}$ , each with equal probability. Write  $S = X_1 + X_2 + \dots + X_{10}$ . Compute the following:

(a) The moment generating function  $M_{X_1}(t)$ . **ANSWER:**

$$M_{X_1}(t) = E[e^{tX_1}] = \frac{1}{6}(e^t + e^{2t} + e^{3t} + e^{4t} + e^{5t} + e^{6t}).$$

(b) The moment generating function  $M_S(t)$ . **ANSWER:** The moment generating function of a sum of independent random variables is the product of the moment generating functions of the individual random variables. Hence  $M_S(t) = \left(\frac{1}{6}(e^t + e^{2t} + e^{3t} + e^{4t} + e^{5t} + e^{6t})\right)^{10}$ .

6. (15 points) Let  $X$  and  $Y$  be random variables with joint density function  $f_{X,Y}(x, y) = \frac{1}{2\pi} e^{-(x^2+y^2)/2}$ . Write  $Z = X + Y$ .

(a) Compute  $E[XY]$ . **ANSWER:**  $X$  and  $Y$  are independent normal random variables, each with mean zero and variance one. Since they are independent we have

$$E[XY] = E[X]E[Y] = 0. \text{ Alternatively, write } E[XY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy \frac{1}{2\pi} e^{-(x^2+y^2)/2} dx dy.$$

Then there are various ways to argue by symmetry that this must be zero.

(b) Compute the conditional expectation  $E[Y|Z]$ . That is, express the random variable  $E[Y|Z]$  in terms of  $Z$ . **ANSWER:** We have  $Z = E[Z|Z] = E[X|Z] + E[Y|Z]$ . Since  $E[X|Z]$  and  $E[Y|Z]$  are the same by symmetry, the answer must be  $Z/2$ .

(c) Compute the probability  $P(X^2 + Y^2 \leq 4)$ . **ANSWER:** This can be computed using polar coordinates. The integral becomes

$$\int_0^2 \int_0^{2\pi} \frac{1}{2\pi} e^{-r^2/2} r d\theta dr = \int_0^2 e^{-r^2/2} r dr = -e^{-r^2/2} \Big|_0^2 = -e^{-2} - (-1) = 1 - e^{-2} \approx .86466$$

MIT OpenCourseWare  
<https://ocw.mit.edu>

18.600 Probability and Random Variables  
Fall 2019

For information about citing these materials or our Terms of Use, visit: <https://ocw.mit.edu/terms>.